

## 11. Hastings-Metropolis Algorithm – Lecture 11 – October 8, 2002

### 11.1 The Hastings-Metropolis Algorithm

1. Begin with an irreducible Markov matrix  $Q_{ij}$ , with  $i, j = 1, 2, 3, \dots, n$  (which need not be symmetric).
2. Let  $n = 0$  and  $X_0 = k$ , for some  $1 \leq k \leq m$ . *what is m?*
3. Generate a new random variable  $X$  such that  $\text{Prob}\{X = j\} = Q_{X_n, j}$ .
4. Generate a random number  $U$  uniformly distributed on  $(0, 1)$ . If  $U < \frac{[b(X)Q_{X, X_n}]}{[b(X_n)Q_{X_n, X}]}$  then  $NS = X$ ; otherwise  $NS = X_n$ .
5. Let  $n = n + 1$ ; set  $X_n = NS$ . *what is NS?*
6. Go to step 3.

### 11.2 Application – Example 10a in Ross

We begin with a large set  $L$  of all permutations of  $\{1, 2, \dots, n\}$  for which  $\sum_{j=1}^n jx_j > a$  for a given constant  $a$ . Another example is the set of all tree subgraphs of a given graph  $H$ . We want a limiting probability distribution which is uniform.

– First define a Markov chain graph  $G$  whose vertices are the elements of  $L$ . We will need a notion of neighbors in  $L$ , and we join node  $i$  to  $j$  by an arc if  $j$  is accessible from  $i$  in one move.

– For  $L = S_n$  we put the arcs of  $G$  between states or permutations which differ by a transposition.

– Let  $N(s) = \{\text{the neighbors of a node } s\}$ , and let  $|N(x)|$  equal the cardinality of  $N(s)$ . Let  $Q_{s,t} = \frac{1}{|N(s)|}$  if  $t \in N(s)$ .

– Since we are interested in sampling uniformly from  $L$ , we want  $\Pi(s) = \Pi(t) = K = \frac{b_s}{\sum b_j}$ . Therefore, by setting

$$\min(1, \frac{b_t Q_{ts}}{b_s Q_{st}}) = \min(\frac{|N_s|}{|N_t|}, 1) \quad (11.1)$$

we get an ergodic Markov chain which is reversible:

$$P_{s,t} = \begin{cases} Q_{s,t} \min(1, \frac{b_t Q_{ts}}{b_s Q_{st}}) \\ Q_{s,s} + \sum_{r \neq s} Q_{s,r} (1 - \min(1, \frac{b_r Q_{rs}}{b_s Q_{sr}})) \end{cases} \quad (11.2)$$

by a theorem in the last lecture.

– The justification for using  $\min(1, \frac{b_t Q_{ts}}{b_s Q_{st}}) = \min(1, \frac{|N_s|}{|N_t|})$  as opposed to the earlier formulation  $\min(1, \frac{b_s}{b_t})$  is based on relaxing the condition that  $Q_{st} = Q_{ts}$ .

We check that  $\sum_{t=1}^m P_{st} = 1$ ; and  $P_{ss} > 0$  for some  $s$ , which is a generic property that leads to aperiodicity.

– Suppose the current state is  $X_n = S$ . Choose one of its neighbors randomly (make a random transposition).

If  $|N_t| \leq |N_s|$ , then accept  $X_{n+1} = t$ .

If  $|N_t| > |N_s|$ , then set  $X_{n+1} = t$  with probability  $q = \frac{|N_s|}{|N_t|}$ . Set  $X_{n+1} = s$  with probability  $1 - q = 1 - \frac{|N_s|}{|N_t|}$ .

### 11.3 Gibbs Sampler (Vector form of Hastings-Metropolis)

Let  $\vec{X} = (X_1, x_2, \dots, x_n)$  be a random vector with probability mass function  $p(\vec{X}) = cg(\vec{x})$ . We want to sample from such a distribution of random vectors, where  $g(\vec{X})$  is known but  $c$  is not.

Consider a Markov chain where the states are  $\vec{x} = (x_1, x_2, \dots, x_n)$ . Let  $\vec{x}$  be the current vector state. Choose  $i = 1, 2, \dots, n$  randomly and set the random variable  $X = x$  with:

$$Prob\{X = x\} = Prob\{X_i = x | X_j = x_j, j \neq i\} \quad (11.3)$$

(which is given a priori by  $p(\vec{x})$ ), since

$$Prob\{X_i = x | X_j = x_j, j \neq i\} = \frac{Prob\{X_i = x, X_j = x_j, j \neq i\}}{Prob\{X_j = x_j, j \neq i\}} \quad (11.4)$$

$$= \frac{Prob\{X_i = x, X_j = x_j, j \neq i\}}{\sum_k Prob\{X_i = k, X_j = x_j, j \neq i\}} \quad (11.5)$$

$$= \frac{p(x_1, x_2, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n)}{\sum_k p(x_1, x_2, \dots, x_{i-1}, k, x_{i+1}, \dots, x_n)} \quad (11.6)$$

Next, if the random variable  $X = x$ , then  $\vec{y} = (x_1, x_2, \dots, x_{i-1}, x, x_{i+1}, \dots, x_n)$  is a possible next state vector.

This is equivalent to Hastings-Metropolis with:

$$Q(\vec{x}, \vec{y}) = \frac{1}{n} \text{Prob}\{X_i = x | X_j = x_j, j \neq i\} \quad (11.7)$$

$$= \frac{p(\vec{y})}{n \text{Prob}\{X_j = x_j, j \neq i\}} \quad (11.8)$$

$$= \frac{p(\vec{y})}{n \sum_k p(x_1, x_2, \dots, x_{i-1}, k, x_{i+1}, \dots, x_n)} \quad (11.9)$$

and

$$P_{\vec{x}, \vec{y}} = \left\{ \begin{array}{l} Q(\vec{x}, \vec{y}) \min(1, \frac{p(\vec{y})Q(\vec{y}, \vec{x})}{p(\vec{x})Q(\vec{x}, \vec{y})}) \\ Q(\vec{x}, \vec{x}) + \sum_{\vec{z} \neq \vec{x}} Q(\vec{x}, \vec{z}) (1 - \frac{p(\vec{z})Q(\vec{z}, \vec{x})}{p(\vec{x})Q(\vec{x}, \vec{z})}) \end{array} \right\} \quad (11.10)$$

$$= \left\{ \begin{array}{ll} Q(\vec{x}, \vec{y}) & \vec{y} \neq \vec{x} \\ Q(\vec{x}, \vec{x}) & \vec{y} = \vec{x} \end{array} \right\} \quad (11.11)$$

## 11.4 Application (Example 10b)

the problem is to generate  $n$  random points on the unit circle, such that no two points are within distance  $d$  of each other, where

$$\beta = \text{Prob}\{\text{no two points are within distance } d \text{ of each other}\} \quad (11.12)$$

is assumed to be small.

Do this by applying the Gibbs sampler, starting with  $n$  points on the unit sphere  $x_1, x_2, \dots, x_n$  such that no two are within distance  $d$  of each other.

Generate a random number  $U$  and let  $I = \text{int}(nU) + 1$ . this step picks randomly from  $i = 1, 2, \dots, n$ .

Next, generate a random point on the circle,  $X = x$  and if  $|x - x_j| > d, j \neq I$ , then set  $\vec{y} = (x_1, x_2, \dots, x_{I-1}, x, x_{I+1}, \dots, x_n)$ . Otherwise generate another point  $X = x'$  and repeat.

## 11.5 The Metropolis Algorithm for Statistical Mechanics

Construct Aafinal probability distribution

$$\Pi_{\vec{x}} = P_{\vec{x}} = \frac{e^{-\beta E(\vec{x})}}{Z_N} \equiv \frac{b(\vec{x})}{\sum_{\vec{r}} b(\vec{r})} \quad (11.13)$$

where  $N$  denotes the number of particles or the number of lattice sites. To be precise let us suppose that the domain is the unit square (with periodic boundary conditions), and the  $N$  particles can take any position on a uniform ( $M \times M$ ) grid – let us say (100x100) for this demonstration.

a. Compute the number of states in the problem. For  $N = 100$  this number is  $(10^4)^M = (10^4)^{100}$  which is huge!

b. So  $Q$  is an  $M^{2N} \times M^{2N}$  matrix  $Q(\vec{x}, \vec{y}) = \frac{1}{N} \text{Prob}\{X_i = x | X_j = x_j, j \neq i\}$ .

c. The important point is that in the Metropolis algorithm,

$$P_{\vec{x}, \vec{y}} = \left\{ \begin{array}{l} Q(\vec{x}, \vec{y}) \min(1, \frac{b_{\vec{y}} Q(\vec{y}, \vec{x})}{b_{\vec{x}} Q(\vec{x}, \vec{y})}) \\ Q(\vec{x}, \vec{x}) + \sum_{\vec{z} \neq \vec{x}, b_j < b_i} Q(\vec{x}, \vec{z}) (1 - \frac{b_{\vec{y}} Q(\vec{y}, \vec{x})}{b_{\vec{x}} Q(\vec{x}, \vec{y})}) \end{array} \right\} \quad (11.14)$$

d. Finally, we need to analyze what

$$\text{Prob}\{X_i = x | X_j = x_j, j \neq i\}; i, j = 1, 2, \dots, N \quad (11.15)$$

$x_j \in L$ , the  $M \times M$  grid.

This can be treated as a graph  $G$  where the nodes are all possible states  $\vec{x}$  of which there are  $M^{2N}$ . Arcs connect “neighbors” in  $G$  of  $\vec{x}$  which are defined by states  $\vec{y}$  that can be reached from  $\vec{x}$  by a Gibbs sampler move, i.e, randomly choose  $i = 1, 2, \dots, N$  and then change  $X_i = x$ , leaving  $X_j = x_j$  for  $j \neq i$  with

$$\text{Prob} X_i = x | X_j = x_j = \frac{\text{Prob}\{X_i = x, X_j = x_j\}}{\text{Prob}\{X_j = x_j, j \neq i\}} \quad (11.16)$$

$$= \frac{e^{-\beta E(\vec{y})}}{\sum_k e^{-\beta E(X_i=k)}} \quad (11.17)$$

e. Putting it together we get

$$Q(\vec{x}, \vec{y}) = \frac{e^{-\beta E(\vec{y})}}{N \sum_k e^{-\beta E(X_i=k)}} \quad (11.18)$$

and

$$P_{\vec{x}} = \left\{ \begin{array}{ll} Q(\vec{x}, \vec{y}) & \vec{y} \neq \vec{x} \\ Q(\vec{x}, \vec{x}) & \vec{y} = \vec{x} \end{array} \right\} \quad (11.19)$$